

When Income Effects are Large: Labor Supply Responses and the Value of Welfare Transfers

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VisitINPS Web Conference
July 9, 2020

The findings and conclusions expressed are solely those of the author and do not represent the views of INPS.

Motivation

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- Yet, **hard to come by** → income effects assumed away or calibrated

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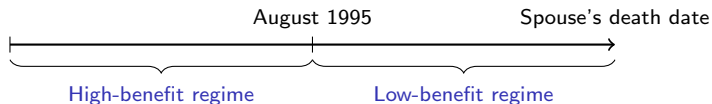
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- Still know surprisingly little about income effects of welfare transfers, mostly due to identification challenges
- **Ideal experiment**: random allocation of lump sums [Cesarini et al., 2017]
- Yet, **hard to come by** → income effects assumed away or calibrated
- Quasi-experimental evidence
 - Based on **short-lived, modest, anticipated** transfers → attenuation
 - Finds overall **small income effects** on labor supply (≈ -0.10)

This paper

- New estimates of **long-term income effect** of welfare transfers on *(i)* labor supply, *(ii)* earnings and *(iii)* total income

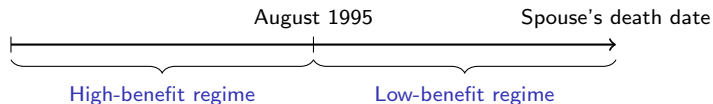
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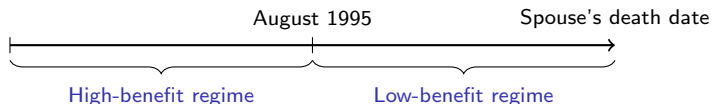
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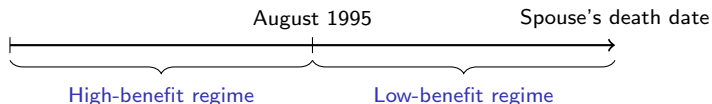
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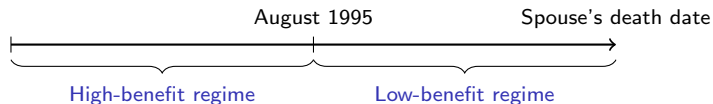
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- **Strategy**: compare **long-term outcomes** of otherwise identical individuals receiving high vs. low benefits for rest of their lives
- **Data**: new **admin data** on universe of benefits and working histories
- **Size**: expected lifetime benefit $\downarrow \approx$ €100,000 or €2,000 per year \triangleright **Large!**

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- **Unique policy experiment** in Italy: **large** and **permanent** reduction in survivor insurance benefits



- **Strategy**: compare **long-term outcomes** of otherwise identical individuals receiving high vs. low benefits for rest of their lives
 - ▷ Unique window on long-run response to benefit change
 - ▷ Comparison of widow(er)s with widow(er)s

Key insights

1. Long-run **income effect** of benefit on earned income $\approx -1 \rightarrow$ **Large!**
 - Fully driven by **extensive** margin (\uparrow entry and \downarrow retirement)
 - No effect on **intensive** margin nor on **wage** rate
 - Large **program substitution** responses (paid family leave, UI)
 - Dynamics of LFP consistent with short-run optimization **frictions**

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 - ii. **High value** of marginal \$ of transfer?

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Revealed-preference model to infer value of benefit from LFP response

Marginal \$ 50% more valuable in low- vs high-benefit regime

Related literature and main contribution

Income effects on labor supply

- NIT (Robins, 1985; Burtless, 1986; Ashenfelter and Plant, 1990; Hum and Simpson, 1993), SSI (Deshpande, 2016), social security (Gelber et al, 2016)
- Unconditional cash transfers (Akee et al., 2010; Jones and Marinescu, 2017)
- Lottery wins (Imbens et al., 2001; Cesarini et al., 2017)

Micro vs. macro elasticities: optimization frictions and indivisibility of labor

- Chetty (2012), Ljungqvist and Sargent (2007), Rogerson and Wallenius (2009)

Optimization methods to measure value of insurance

- UI (Chetty, 2008; Landais, 2015; Hendren, 2017; Landais and Spinnewijn, 2019), health shocks (Fadlon and Nielsen, 2018; Dobkin et al., 2018)

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Optimization methods to measure value of insurance

- UI (Chetty, 2008; Landais, 2015; Hendren, 2017; Landais and Spinnewijn, 2019), health shocks (Fadlon and Nielsen, 2018; Dobkin et al., 2018)
 - **Revealed-preference method** to quantify **value of transfer**
 - Provide **new estimate** in context of **survivor benefits** (Fadlon et al., 2019)

Outline

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2. Income Effect of Benefit on Taxable and Disposable Income
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Institutional features of Italian survivor benefit program

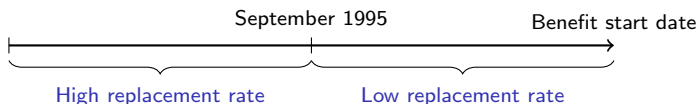
- Social insurance program that provides benefit to relatives of deceased retirees or workers (spouse, dependent children)
 - Focus on surviving spouses
- Benefit (B) is fraction (b) of pension (P) that deceased was entitled to

$$B = b \times P$$

- Benefit starts on first day of calendar month following death event
- Entitlement ends upon remarriage or loss of dependency status

The 1995 reform of survivor benefits

- Reform reduced benefit replacement rate (b) for spouses without dependent children
- Passed on 8 Aug 1995, it applies to all benefit payments starting on or after 1 Sept 1995



- Part of broader reform of social security system (so-called Dini Reform)
→ No confounding effect

Other recipients

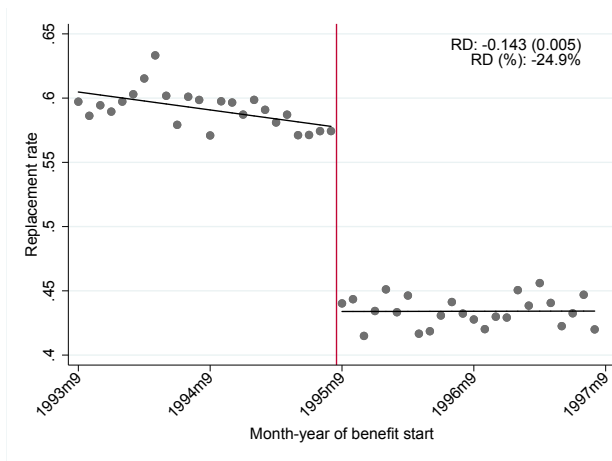
No anticipation

Balancing test

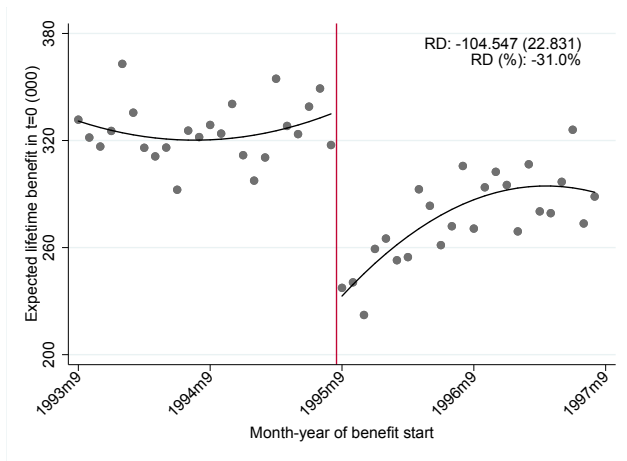
Pension reform

The 1995 reform of survivor benefits

Reduction in benefit replacement rate for spouses without dependent children



Large and permanent reduction in lifetime benefit



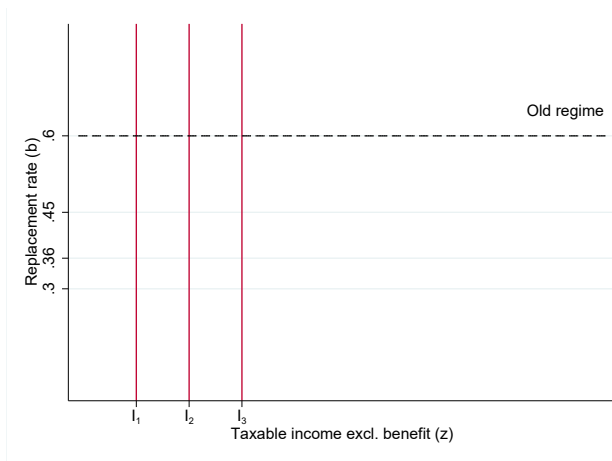
Corresponds to expected **lifetime** benefit drop of **€100,000** [Regressions](#)

Note: Surviving spouses aged 55 and younger at time of spouse's death

[Go to data](#)

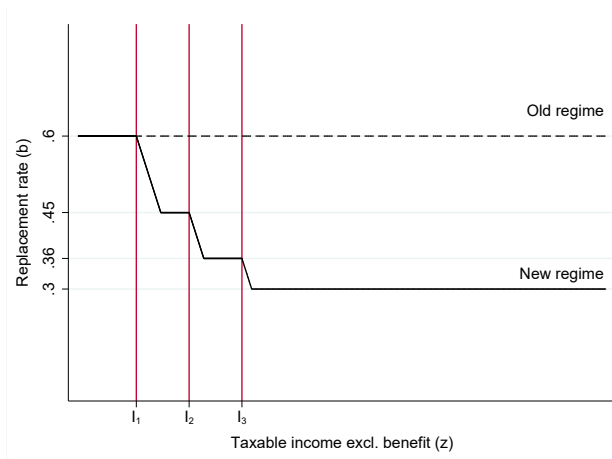
The 1995 reform of survivor benefits

Benefit start date before 1.9.1995



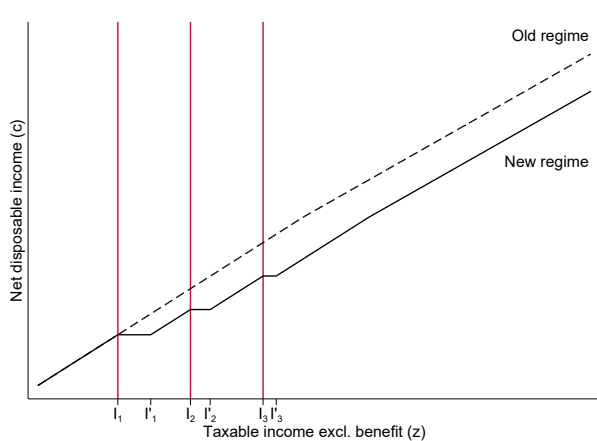
The 1995 reform of survivor benefits

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Minimum pension

Effect of 1995 reform on survivor's static budget set



Personal Income Tax

Dynamic effects

Effects of 1995 reform on survivors' labor supply

Static effects

- Income effect for $z > l_1$
- Substitution effect for $z \in [l_j, l'_j] \rightarrow$ Labor supply \downarrow

Dynamic effects

- Income effect
- Substitution effect: lower net returns to extra year of work \rightarrow Lifetime LS \downarrow

Effects of 1995 reform on survivors' labor supply

Static effects

- Income effect for $z > l_1$
- Substitution effect for $z \in [l_j, l'_j]$ → Labor supply ↓

Dynamic effects

- Income effect
 - Substitution effect: lower net returns to extra year of work → Lifetime LS ↓
- Labor supply response to benefit ↓ is **lower bound** of income effect!

Identification strategy

- Identify effect of B_{it} on Y_{it} using **IV-RD** around 1.9.1995 cutoff

$$Y_{it} = \alpha + \beta \cdot B_{it} + X'_{it} \cdot \gamma + \eta_{it} \quad (1)$$

$$Y_{it} = \alpha_0 + \beta_0 \cdot \mathbb{I}[\tau_i \geq 0] + \sum_{k=1}^K \alpha_k \cdot \tau_i^k + \sum_{k=1}^K \beta_k \cdot \tau_i^k \cdot \mathbb{I}[\tau_i \geq 0] + \varepsilon_{it} \quad (2)$$

τ benefit start date relative to Sept 1995, t time since death RD

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τ benefit start date relative to Sept 1995, t time since death **RD**

- Focus on sample with **predicted** $\mathbf{z} > \mathbf{I}_1 \rightarrow$ income effect **Lasso** **First stage**

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τ benefit start date relative to Sept 1995, t time since death RD

- Focus on sample with **predicted $z > \mathbf{I}_1$** \rightarrow income effect Lasso First stage
- Exploit kinks in budget set to **quantify substitution incentives** via bunching approach

Data

- Administrative data on **universe of survivor benefits** since 1990
 - Start/end date, benefit amount, taxable income, recipients and relationship to deceased
- Linked to **survivors' contributory histories** up to 2017
 - Employment spells, earnings, social insurance take-up, job characteristics
- Linked to demographic archive
- Balanced panel of surviving spouses aged ≤ 55 at time of spouse's death and observed for 15 years after death

Summary statistics

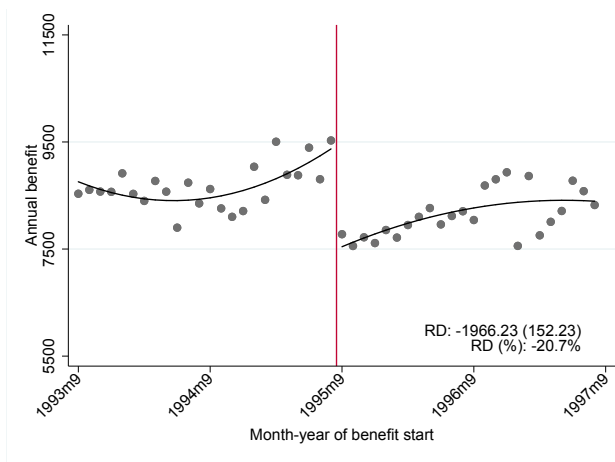
Mortality

Remarriage

Outline

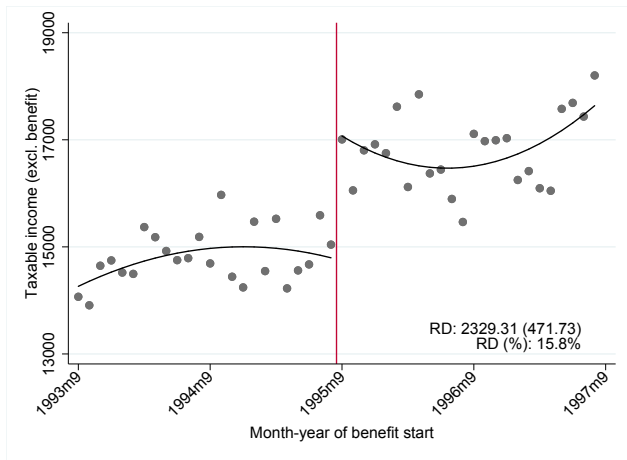
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Average annual drop in benefit amount

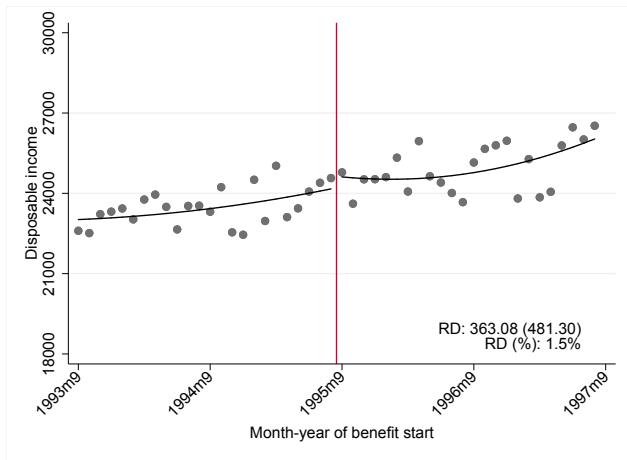


Note: Pooling event time years from $t = 0$ to $t = 15$

Average annual increase in taxable income



No effect on disposable income



Estimation of income effect

- Benefit $\downarrow \approx \text{€}2000$, taxable income $\uparrow \approx \text{€}2000$, disposable income =
- Marginal propensity to earn out of unearned income: $\text{MPE} \equiv \frac{\partial z}{\partial B}$
- IV-RD estimate of MPE ≈ -1

	Taxable income (z) (1)	Disposable income ($z + B$) (2)
Benefit (B)	-1.008*** (0.303)	-0.008 (0.303)
Obs.	216896	216896

- Rescaling by $\frac{\bar{B}}{\bar{z}}$, obtain income elasticity of approx. -0.6

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Mechanisms

Labor supply response

- Large extensive-margin response (\uparrow entry and \downarrow retirement) **LFP**
- Dynamics of LFP reveals short-run optimization frictions **LFP dynamics**
- No intensive margin response **Days**
- No effect on wage rate, not even in long run **Wage**

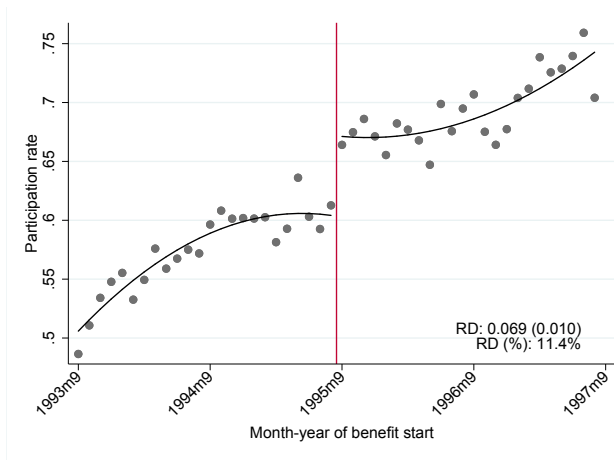
Program substitution

- Increased take-up of paid family leave and UI **Social ins.**
- Extension of dependency period to delay benefit loss **Dependency**

Extensions

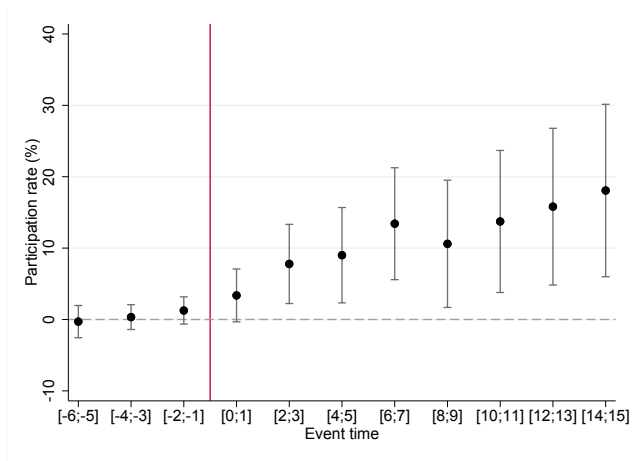
Conclusion

Sharp increase in participation rate



IV-RD ($\times 1000$): -0.044 (0.007) [Placebo](#) [Back](#)

Dynamics of participation response



[In levels](#)

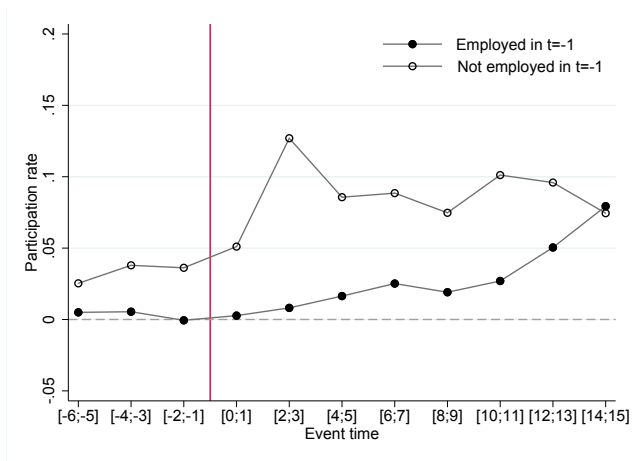
[By gender](#)

[By age in \$t = 0\$](#)

[By dep. children in \$t = 0\$](#)

[Back](#)

Dynamics of participation response: entry and exit



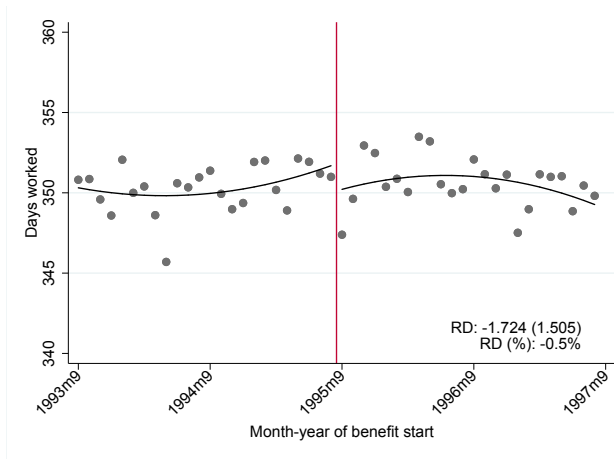
[Decomposition of LFP response](#)

[Retirement rate](#)

[Age profile of entry/exit](#)

[Back](#)

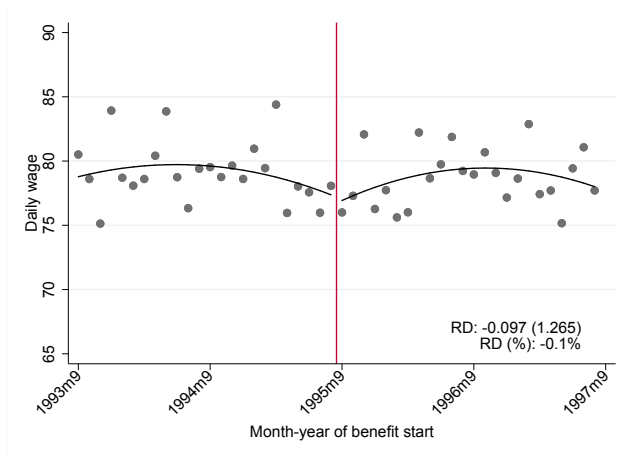
No intensive margin response



Note: Conditional on employment. Individuals employed in $t = -1$.

[Back](#)

No effect on the wage rate



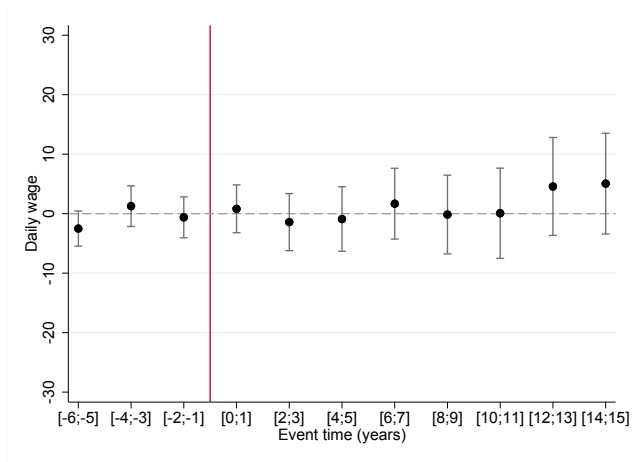
Note: Conditional on employment. Individuals employed in $t = -1$.

Heterogeneity

Other margins of LS adjustment

Back

No long-term effect on wage rate



→ Limited returns to experience, human capital accumulation and effort

[Dynamics of intensive margin](#)

[Back](#)

Program substitution: work-related benefits

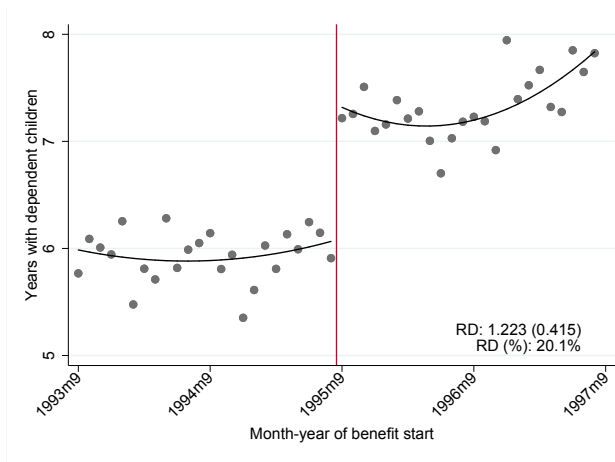
	Paid family leave (1)	Paid sick leave (2)	Unemployment benefits (3)
Benefit ($\times 1000$)	-0.003** (0.001)	-0.001 (0.003)	-0.013*** (0.002)
Obs.	115137	115137	115137
Control mean	0.008	0.042	0.016

Note: Conditional on being employed in t or $t - 1$. Individuals employed in $t = -1$ [Back](#)

Surviving spouses with dependent children

- Benefit replacement rate for spouse with dependent children is higher
 - 80% with 1 child, 100% with 2+ children Replacement rates
- Not affected by 1995 reform
- Incentive for treatment group to extend dependency period Benefit loss
- Dependency status
 - Up to age 18 by default
 - Up to age 26 conditional on high-school/university enrolment

Program substitution: years with dependent children



IV-RD ($\times 1000$): -0.653 (0.272)

Placebo

Timing of LFP response

Back

Extensions

- Normative implications [Detail](#)
- Comparison with existing estimates [Detail](#)
- Compatibility with macro-elasticity [Detail](#)
- Relationship with theoretical models of labor supply [Detail](#)
- External validity and policy relevance [Detail](#)

[Conclusion](#)

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A large income effect: why?

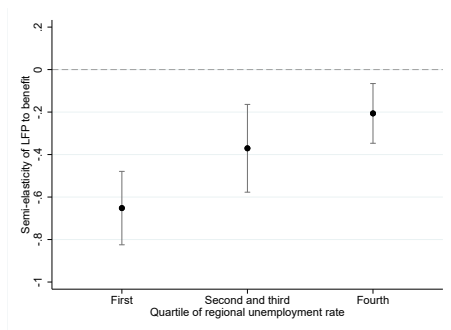
- Benefit losses trigger large labor supply responses
- Potential explanations
 - ▷ **High value** of transfer?
Large income effects reveal highly curved utility, i.e. high risk aversion
 - ▷ **Low cost** of adjusting labor supply?
For given risk aversion, lower adjustment costs imply larger responses
- Understanding which one prevails is important for **welfare** analysis

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Heterogeneity wrt regional unemployment rate

- Cost of adjusting labor supply likely increasing in local unemployment rate (job-search / on-the-job effort cost)
- LFP response to given % \downarrow in B is smaller in regions with higher u rate



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Labor supply responses and the value of transfers

- Agents undertake costly actions to increase c in response to $\downarrow B$
- *Ceteris paribus*, **larger responses** \leftrightarrow **larger implicit valuation** of B
- Value of marginal \$ of transfer captured by

$$MB = \frac{u'(c(0))}{u'(c(B))} - 1$$

- **Revealed-preference** approach maps labor supply response onto MB (Chetty, 2006, 2008)

A revealed-preference approach

- Model of extensive margin labor supply response (LFP rate Φ)
- Participation response (rescaled) reveals value of benefit

$$MB = \frac{\overbrace{-\left(\frac{d\Phi}{d \log B}\right)}^{\text{Income effect}}}{\underbrace{\varepsilon}_{\text{Cost of work (-)}}$$

$\varepsilon = d\Phi/d \log z$ is semi-elasticity of labor supply to labor earnings

Proof

Applicability

Quantifying the value of transfers

- Calibrate using
 - Own estimates of semi-elasticity of LFP to benefit
 - Estimates of $\varepsilon \approx 0.6$ from Blundell et al. (2016)

$$MB = \frac{0.3}{0.6} = 0.5$$

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$$MB = \frac{0.3}{0.6} = 0.5$$

- $u'(c)$ is 50% higher in low- vs high-benefit regime
- Large labor supply responses \rightarrow **widowhood** is state with **high $u'(c)$**

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 - Estimates of $\varepsilon \approx 0.6$ from Blundell et al. (2016)

$$MB = \frac{0.3}{0.6} = 0.5$$

- $u'(c)$ is 50% higher in low- vs high-benefit regime
- Large labor supply responses → **widowhood** is state with **high $u'(c)$**
- Large **welfare gains** from increased survivor insurance generosity

Outline

1. Institutional Context, Identification Strategy and Data
2. Income Effect of Benefit on Taxable and Disposable Income
3. Mechanisms: Labor Supply and Program Substitution
4. Normative Implications of Large Income Effects
 - Evidence on Adjustment Costs: Heterogeneity Analysis
 - Value of Transfers: Revealed-Preference Approach
5. Probing the Large Income Response
 - Comparison w Existing Estimates and Compatibility w Macro-Elasticity
 - Relationship with Theories of Labor Supply
 - External Validity and Policy Relevance

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Comparison with existing estimates

MPE ≈ -1 larger than most existing estimates, with some exceptions

Source	Context	MPE
Robins (1985), Hum and Simpson (1993)	NIT	≥ -0.2
Imbens et al. (2001), Cesarini et al. (2017)	Lottery wins	$= -0.1$
Gelber et al. (2016, 2017)	SS Notch	≤ -0.6 (M), -0.9 (F)
Deshpande (2016)	SSI	≈ -1.4

Potential explanations

- Differences in populations analyzed (risk preferences)
- Asymmetric response to gains/losses (loss aversion, sticky consumption)
- Degree of *ex-ante* insurance and size of income shock
- Fungibility (Thaler, 1990)

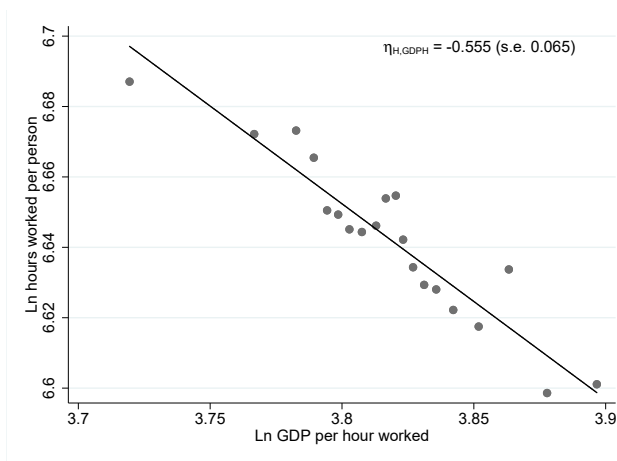
Compatibility with macro-elasticity

- Longstanding debate on discrepancy between micro and macro elasticities
- **Macro** estimates of steady-state elasticities **larger** than **micro** estimates

$$\varepsilon^M \gg \varepsilon^m$$

- Two factors can account for gap:
 - **Frictions** (Chetty, 2012)
 - **Indivisible labor** (Rogerson, 1988; Ljungqvist and Sargent, 2006; Rogerson and Wallenius, 2009)
- Assess compatibility of estimated of long-run micro elasticity ($\hat{\eta} = -0.6$) with macro elasticity

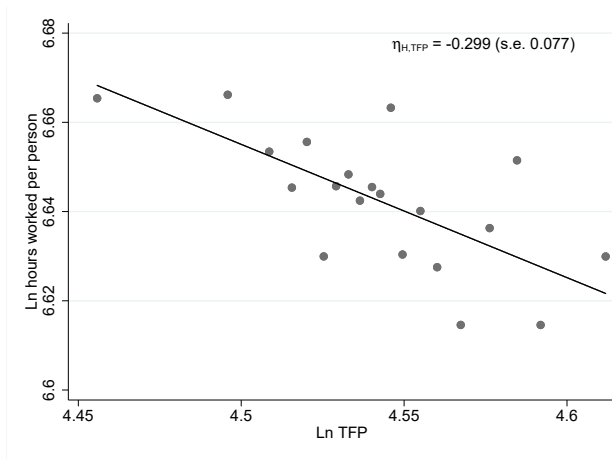
Macro-elasticity of hours per person to GDP per hour



- Data on OECD countries from 1985 to 2015
- Binned scatterplot conditional on country and year fixed effects

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Macro-elasticity of hours per person to TFP



- Data on OECD countries from 1985 to 2015
- Binned scatterplot conditional on country and year fixed effects

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Relationship with theoretical models of labor supply

- Effect of benefit loss on z and c can be rationalized in model with preferences **quasi-linear in work effort**

- $U(c, z) = u(c) - \frac{z}{\theta}$

- $\frac{\partial c^*}{\partial B} = 0$ and $\frac{\partial z^*}{\partial B} = -1$

- **Career-length model** by Ljungqvist and Sargent (2007): dynamic model with predictions consistent with my findings in reduced form
 - Life-cycle model with time-separable preferences and indivisible labor
 - Agent decides what fraction of life to devote to work
 - Model delivers large extensive margin elasticity

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External validity and policy relevance

1. Sample selection

- Retirement behavior and participation responses to pension reforms
- Single parents most at risk of income insecurity and main target of welfare transfers (e.g. EITC)
- Elasticity of labor supply may differ between marriage and widowhood (e.g. leisure complementarities, loneliness, sharing of family duties)

2. Income shock

- Losing spouse at young age is low-probability, unpredictable event
- Households likely limitedly insured against associated income shock
- Larger responses than for predictable shocks (e.g. job loss)

Conclusion

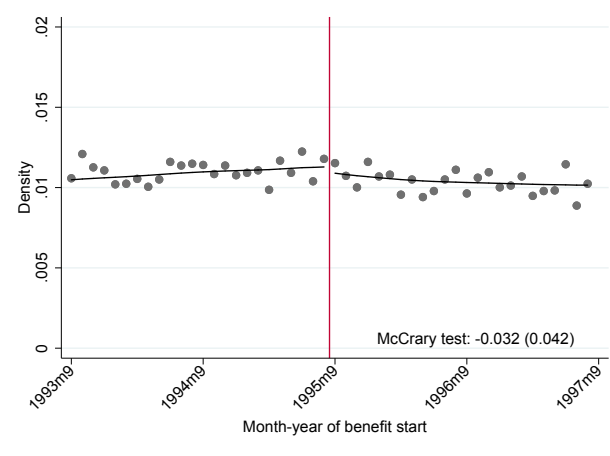
- New evidence on long-run response to permanent benefit drop
 - Combine rich **admin data** and compelling **policy variation**
 - Find **large labor supply response** to benefit loss (MPE ≈ -1)
 - Fully driven by extensive margin
- Examine normative implications
 - Large income effect \rightarrow **large value** of $B \rightarrow$ welfare gains from $\uparrow B$
- Beyond labor supply effects?
 - Intergenerational outcomes
 - General eq. effects on human capital, marriage, fertility choices
 - Perception of welfare state

Appendix

Benefit replacement rate

	Benefit start date	
	Before Sept 1, 1995	After Sept 1, 1995
<i>Spouse (with and without children)</i>		
Spouse only		
Survivor's taxable income $\leq 3 \times$ minimum pension	60%	60%
Survivor's taxable income $\leq 4 \times$ minimum pension	60%	45%
Survivor's taxable income $\leq 5 \times$ minimum pension	60%	36%
Survivor's taxable income $> 5 \times$ minimum pension	60%	30%
Spouse with one child	80%	80%
Spouse with two or more children	100%	100%
<i>Children (absent the spouse)</i>		
One child	60%	70%
Two children	80%	80%
Three or more children	100%	100%
<i>Parents or siblings (absent the spouse, children or grandchildren)</i>		
Each relative	15%	15%

Density test



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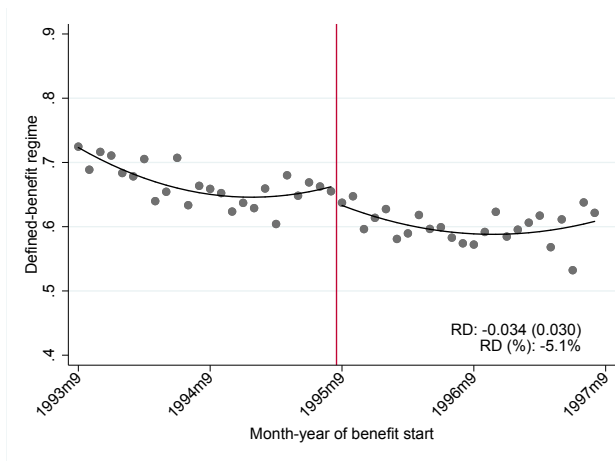
Balancing test

	Regression discontinuity					Control mean
	(1)	(2)	(3)	(4)	(5)	
Female	0.003 (0.004)	0.005 (0.005)	0.006 (0.006)	-0.001 (0.007)	0.003 (0.007)	0.899
Age in $t = 0$	0.070 (0.094)	-0.097 (0.124)	-0.160 (0.143)	-0.216 (0.179)	-0.075 (0.120)	46.860
Experience in $t = -1$	-0.001 (0.143)	-0.006 (0.189)	-0.418* (0.216)	-0.188 (0.269)	-0.289 (0.123)	14.445
Earnings in $t = -1$	-269.993* (139.426)	-170.699 (185.312)	-140.946 (211.708)	-105.928 (265.219)	-111.621 (224.762)	6373.42
Prop. employed in $t = -1$	0.002 (0.006)	-0.005 (0.008)	-0.002 (0.010)	-0.005 (0.012)	0.004 (0.011)	0.397
Days worked in $t = -1$	-0.999 (1.358)	-1.181 (1.797)	-3.594* (2.007)	-0.214 (2.539)	0.165 (3.165)	341.026
Daily wage in $t = -1$	-1.282 (1.020)	1.675 (1.721)	1.189 (1.353)	-0.953 (2.469)	-3.376 (2.204)	47.544
Prop. on defined benefit	-0.005 (0.007)	-0.013 (0.009)	-0.006 (0.010)	-0.011 (0.012)	-0.003 (0.009)	0.312
Obs.	94578	94578	94578	94578	94578	-
Month-of-benefit-start FE		x		x		-
Calendar year FE		x		x		-
Linear trend	x	x	x	x		-
Quadratic trend			x	x		-
LLR					x	-

Dini pension reform (Law 335/1995)

- Pension system transitioned from defined benefit to notionally defined contribution
 - ≥ 18 years of contribution on 31.12.1995 \rightarrow DB
 - < 18 years of contribution on 31.12.1995 \rightarrow pro-rata DB/NDC
 - Starting to contribute on or after 1.1.1996 \rightarrow NDC
- Financing remains pay-as-you-go

Proportion under defined-benefit system



Back

First stage in $t = 0$

	Regression discontinuity				Control mean
	(1)	(2)	(3)	(4)	(5)
Benefit in $t = 0$	-1510.21*** (260.413)	-1684.83*** (296.800)	-2137.66*** (376.689)	-1963.66*** (407.618)	8494.83
Lifetime benefit, 000	-67.032*** (13.811)	-85.273*** (16.691)	-99.641*** (20.155)	-104.547*** (22.831)	337.387
Obs.	13556	13556	13556	13556	-
Benefit-start-month FE		x		x	-
Calendar year FE		x		x	-
Linear trend	x	x	x	x	-
Quadratic trend			x	x	-

Back

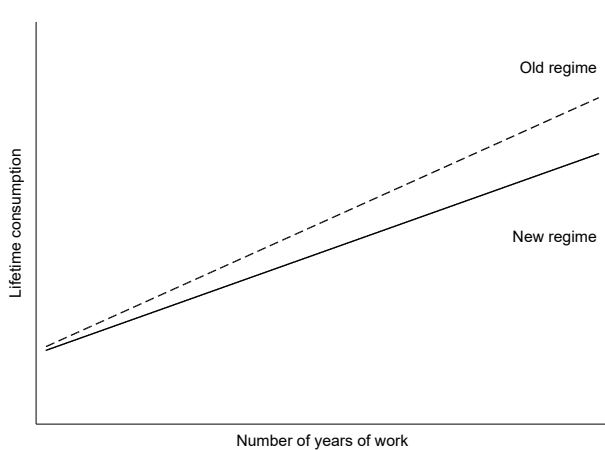
Annual minimum pension (EUR)

Year	Amount	× 3	× 4	× 5
1995	4,205.95	12,617.84	16,823.79	21,029.74
1996	4,433.21	13,299.64	17,732.86	22,166.07
1997	4,606.10	13,818.29	18,424.39	23,030.49
1998	4,684.32	14,052.95	18,737.26	23,421.58
1999	4,768.58	14,305.73	19,074.30	23,842.88
2000	4,844.78	14,534.34	19,379.12	24,223.89
2001	4,970.67	14,912.00	19,882.66	24,853.33
2002	5,104.97	15,314.91	20,419.88	25,524.85
2003	5,227.56	15,682.68	20,910.24	26,137.80
2004	5,358.34	16,075.02	21,433.36	26,791.70
2005	5,465.59	16,396.77	21,862.36	27,327.95
2006	5,558.54	16,675.62	22,234.16	27,792.70
2007	5,669.82	17,009.46	22,679.28	28,349.10
2008	5,760.56	17,281.68	23,042.24	28,802.80
2009	5,950.88	17,852.64	23,803.52	29,754.40
2010	5,992.61	17,977.83	23,970.44	29,963.05

Interaction with personal income tax (PIT)

- z is individual taxable income, including income from labor, retirement, pensions, capital and rents, and excluding survivor benefit B
- Both z and B are subject to personal income taxation
- Disposable income is $c = z + B - T(z + B)$
- Both new and old regimes subject to same PIT schedule
 - Income taxes may add to the wedge between old- and new-regime survivors \rightarrow positive effect on LS (upper bound)
 - PIT brackets do not coincide with survivor benefit brackets

Dynamic effects of 1995 reform



Back

Parametric RD specification

$$Y_{it} = \alpha_0 + \beta_0 \cdot \mathbb{I}[\tau_i \geq 0] + \sum_{k=1}^K \alpha_k \cdot \tau_i^k + \sum_{k=1}^K \beta_k \cdot \tau_i^k \cdot \mathbb{I}[\tau_i \geq 0] + \varepsilon_{i\tau}$$

- Y_{it} is outcome Y for individual i at event time t
- τ_i is benefit start date normalized so that $\tau = 0$ at 1.9.1995 cutoff
- Based on balancing tests, include polynomials in τ of order $K = 2$
- Coefficient of interest capturing effect at cutoff is β_0
- Estimates based on month-of-benefit-start bins and symmetric bandwidth of 24 months (MSE optimal for benefit amount)

Predicting counterfactual income bracket

- Select 10 percent of individuals in the control group
- Predict income bracket at event time $t = 10$ using rich set of observables
 - Predictors include demographics and working history in $t < 0$
 - Use Lasso estimator to select subset of most relevant predictors
- Apply estimated coefficients to treatment group

Back

First stage in $t = 0$

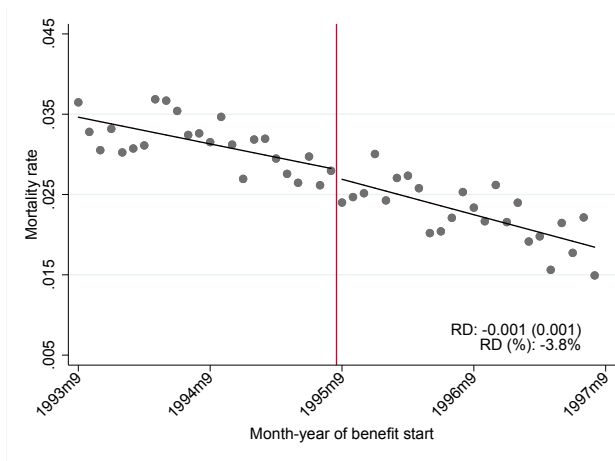
	(1)	Regression discontinuity		(4)	Control mean (5)
		(2)	(3)		
<i>Predicted second or higher income bracket</i>					
Benefit in $t = 0$	-1510.21*** (260.413)	-1684.83*** (296.800)	-2137.66*** (376.689)	-1963.66*** (407.618)	8494.83
Lifetime benefit (000)	-67.032*** (13.811)	-85.273*** (16.691)	-99.641*** (20.155)	-104.547*** (22.831)	337.387
Obs.	13556	13556	13556	13556	-
<i>Full sample</i>					
Benefit in $t = 0$	-465.171*** (73.548)	-558.993*** (85.830)	-593.922*** (109.989)	-602.776*** (120.938)	8371.92
Lifetime benefit (000)	-18.917*** (3.243)	-24.567*** (3.916)	-23.623*** (4.879)	-25.120*** (5.511)	298.57
Observations	94578	94578	94578	94578	-

Summary statistics

	Full sample		Treatment group		Control group	
	Mean	St. Dev.	Mean	St. Dev.	Mean	St. Dev.
Female	0.64	0.48	0.66	0.48	0.62	0.48
Age in $t = 0$	43.50	7.49	43.56	7.31	43.45	7.65
Prop. aged < 40 in $t = 0$	0.29	0.45	0.28	0.45	0.30	0.46
Prop. aged 40-50 in $t = 0$	0.51	0.50	0.53	0.50	0.49	0.50
Prop. aged 51-59 in $t = 0$	0.20	0.40	0.19	0.39	0.21	0.41
Prop. with dependent children in $t = 0$	0.58	0.49	0.58	0.49	0.59	0.49
Age of dependent children in $t = 0$	12.23	5.61	12.29	5.61	12.18	5.62
Prop. ever employed in $t \leq -1$	1.00	0.05	1.00	0.04	1.00	0.06
Years of experience in $t = -1$	20.81	8.85	20.83	8.75	20.78	8.94
Prop. employed in $t = -1$	0.96	0.19	0.96	0.18	0.96	0.19
Prop. empl. in private sector in $t = -1$	0.61	0.49	0.60	0.49	0.62	0.48
Prop. empl. in public sector in $t = -1$	0.14	0.35	0.15	0.36	0.14	0.34
Prop. self-employed in $t = -1$	0.17	0.38	0.17	0.38	0.17	0.38
Labor income in $t = -1$	24216.42	12681.93	24096.99	12625.93	24328.48	12734.13
Daily wage in $t = -1$	72.36	40.08	71.86	38.14	72.82	41.82
Days worked in $t = -1$	347.55	53.53	346.83	55.00	348.22	52.10
Benefit in $t = 0$	762.18	712.46	745.52	686.12	778.00	736.29
Income of deceased in $t = 0$	21361.10	21933.74	21886.54	20968.13	20589.71	23261.99
Pension of deceased in $t = 0$	14104.45	13660.71	14528.82	12980.38	13701.51	14265.97
Observations	13556		6562		6994	

Note: (i) 24-month bandwidth, (ii) monetary quantities in 2010 prices [Back](#)

Mortality



Note: Probability of having died by $t = 15$. Full sample [Back](#)

Remarriage

	Regression discontinuity				Control mean
	(1)	(2)	(3)	(4)	(5)
Remarriage rate ($t = 15$)	-0.002 (0.015)	0.005 (0.017)	0.026 (0.022)	0.045* (0.024)	0.056
Observations	13556	13556	13556	13556	-
Time to remarriage	-0.344 (0.862)	-0.498 (1.050)	-0.627 (1.298)	-2.540* (1.498)	10.116
Observations	1073	1073	1073	1073	-
Benefit-start-month FE		x		x	-
Calendar year FE		x		x	-
Linear trend	x	x	x	x	-
Quadratic trend			x	x	-

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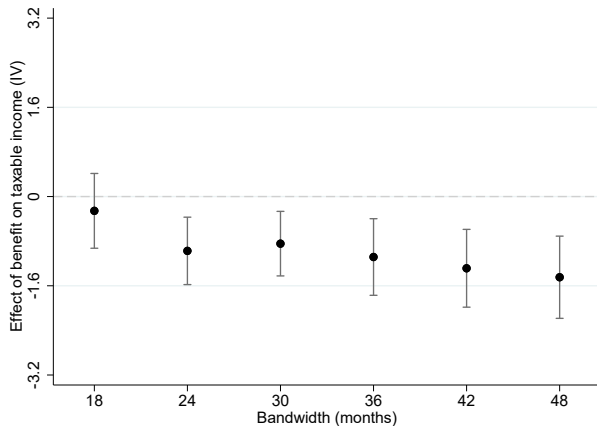
Robustness to parametric specification

	Regression discontinuity				Control mean
	(1)	(2)	(3)	(4)	(5)
Benefit	-1155.25*** (103.033)	-1306.96*** (110.320)	-1771.21*** (145.140)	-1966.23*** (152.225)	9462.31
Taxable income	1674.92*** (380.664)	1473.23*** (407.731)	2508.59*** (455.254)	2329.31*** (471.733)	14470.64
Disposable income	519.674 (386.337)	166.277 (414.151)	737.385 (464.363)	363.081 (481.298)	23932.95
MPE		-1.205*** (0.337)		-1.008*** (0.303)	
Observations	216896	216896	216896	216896	-
Benefit-start-month FE		x		x	-
Calendar year FE		x		x	-
Linear trend	x	x	x	x	-
Quadratic trend			x	x	-

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Robustness to bandwidth choice

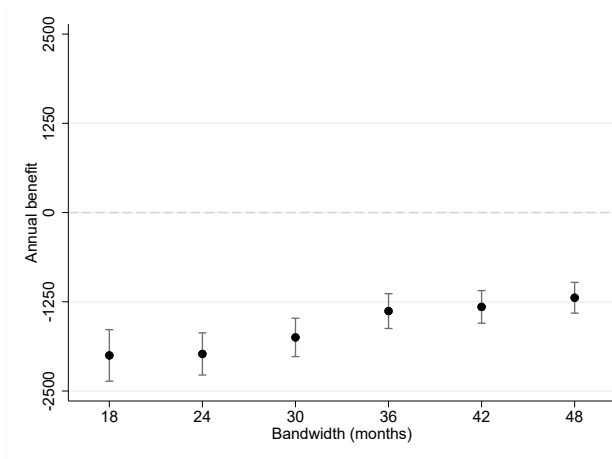
IV-RD on taxable income



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Robustness to bandwidth choice

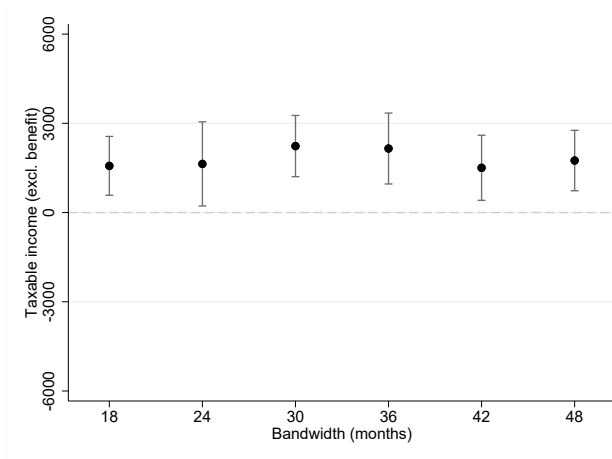
Annual benefit



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Robustness to bandwidth choice

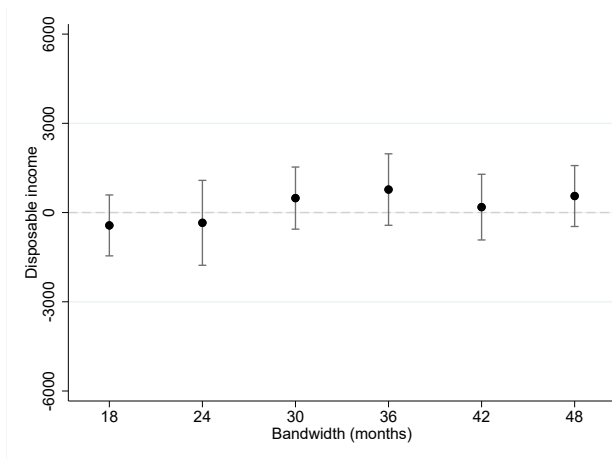
Taxable income



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Robustness to bandwidth choice

Disposable income



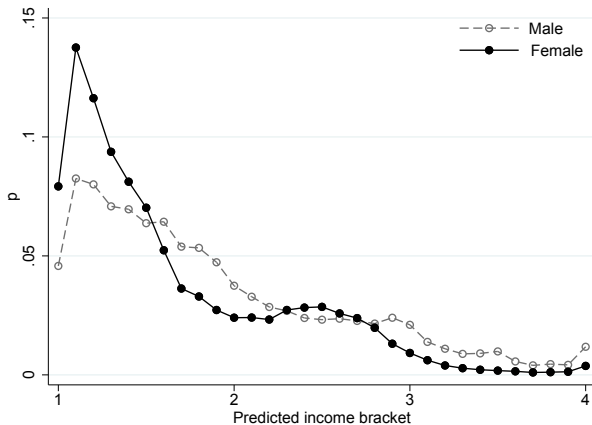
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Heterogeneity by gender and age in $t = 0$

	Gender		Age in $t = 0$		
	Female (1)	Male (2)	20-40 (3)	41-50 (4)	51-55 (5)
Benefit	-1984.11*** (208.525) [11318.84]	-734.45*** (89.437) [7129.74]	-2840.77*** (174.841) [8842.95]	-1194.09*** (245.582) [9612.85]	-2099.00*** (294.558) [8944.35]
MPE	-1.325*** (0.376)	-0.106 (0.772)	-1.097*** (0.459)	-0.999 (0.644)	-0.451 (0.299)
Month FE	x	x	x	x	x
Year FE	x	x	x	x	x
Linear tr.	x	x	x	x	x
Quadratic tr.	x	x	x	x	x

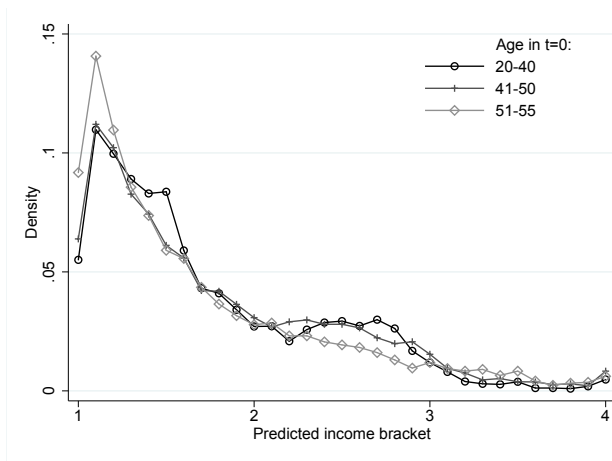
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Empirical density of predicted bracket by gender



Back

Empirical density of predicted bracket by age in $t = 0$



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Heterogeneity by presence of dependent children in $t = 0$

	With dependent children (1)	Without dependent children (2)
Benefit	-779.081*** (155.037) [10035.68]	-2039.88*** (159.371) [8128.27]
MPE	0.636 (0.617)	-1.757*** (0.322)
Month FE	x	x
Year FE	x	x
Linear tr.	x	x
Quadratic tr.	x	x

From reduced-form evidence to income effects

If **substitution incentives** matter, then

- Compensated elasticity $\varepsilon^c > 0$
- Previous estimate is **lower bound** of true income effect

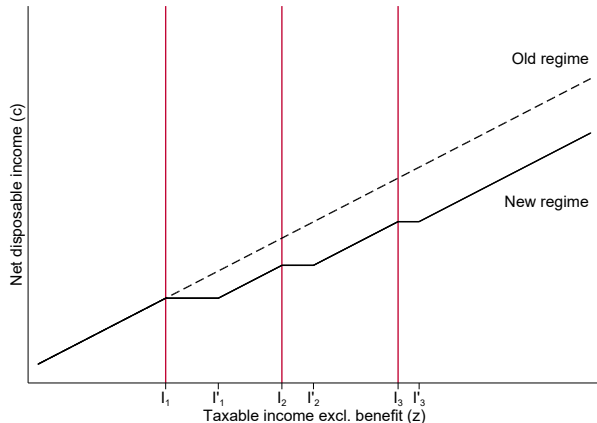
How important are substitution incentives?

1. Estimate based on sample with predicted income in affected range
 - Robust to excluding observations around kinks [Excluding obs. kinks](#)
2. Exploit variation in marginal tax rate at convex kinks to quantify ε^c
 - No bunching at convex kinks consistent with $\varepsilon^c = 0$

[Bunching](#)

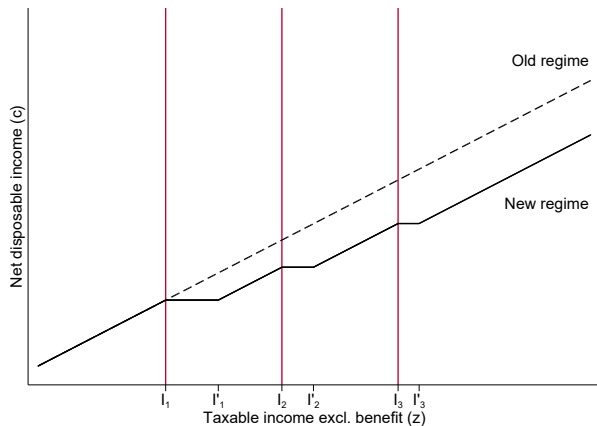
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Reform creates three convex kinks in budget set



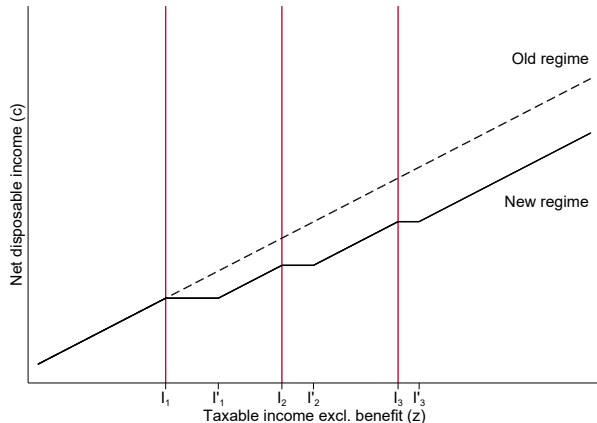
- Reform creates discontinuity in marginal tax rate at $z = l_j$

Reform creates three convex kinks in budget set



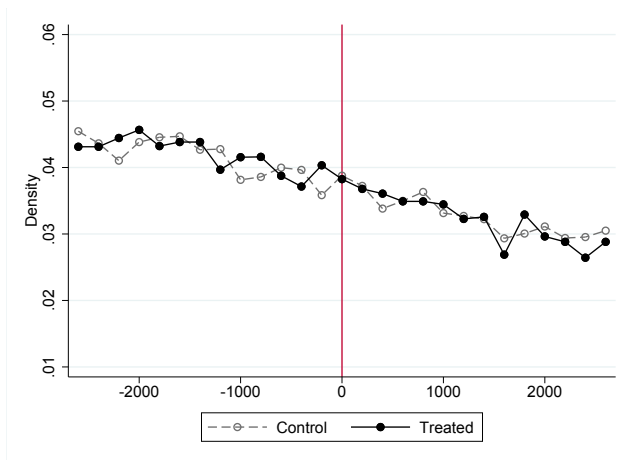
- Treated individuals with $z \in [l_j, l'_j]$ have incentive to bunch at l_j

Reform creates three convex kinks in budget set



- Amount of excess bunching at l_j is proportional to ε^c (Saez, 2010)

No long-run excess bunching at convex kinks



- Not even among self-employed Self-employment
- Consistent with $\varepsilon^c = 0$

Excluding observations around kinks

	Taxable income (1)	Disposable income (2)	Taxable income (3)	Disposable income (4)
Benefit	-0.943** (0.450)	0.057 (0.450)	-0.847** (0.419)	0.153 (0.419)
Observations	73783	73783	73783	73783
Linear trend	x	x	x	x
Quadratic trend			x	x

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Bunching approach

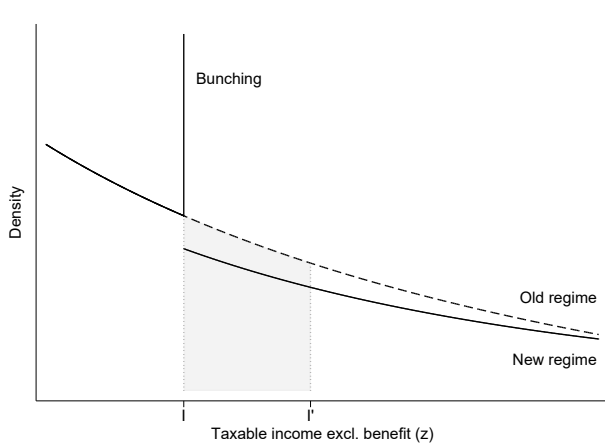
- Let utility be defined over disposable income and taxable income

$$U = u(z - T(z), z/\theta)$$

where $\theta \sim F(\theta)$ is income generating ability and $T(\cdot)$ tax/benefit schedule

- Linear $T(\cdot)$ + smooth $F(\theta) \Rightarrow$ distribution of z smooth
- Reform introduces discontinuity in marginal tax rate at $z = l$ creating *convex kink* in budget constraint of treatment group
- Treated individuals in $[l, l']$ have incentive to bunch at l
 - Excess bunching at l
 - Left-shift of density above l
- Amount of excess bunching in l is proportional to ε^c (Saez, 2010)

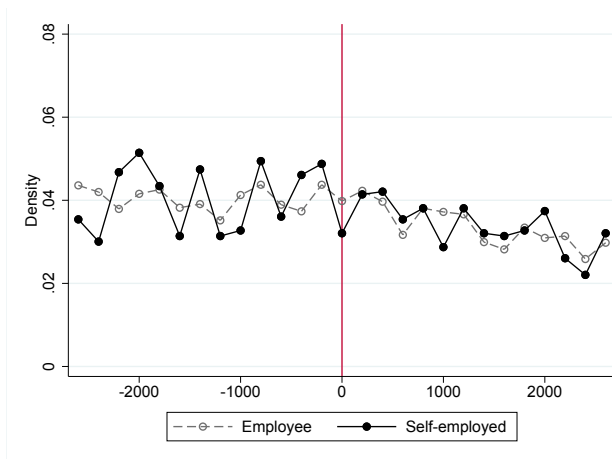
Density of taxable income z



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Empirical density of taxable income by employment status

Treatment group



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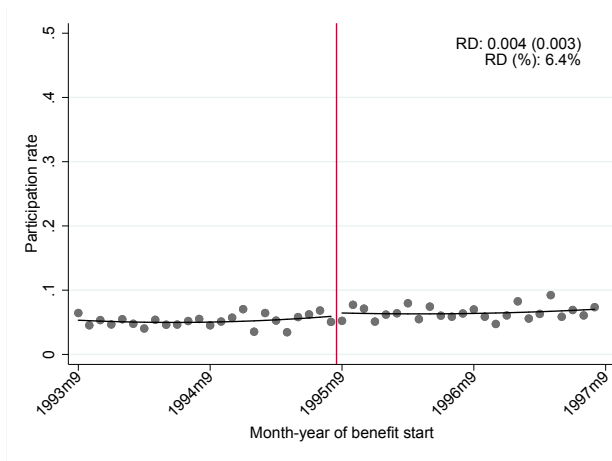
Empirical density of taxable income by employment status

Control group



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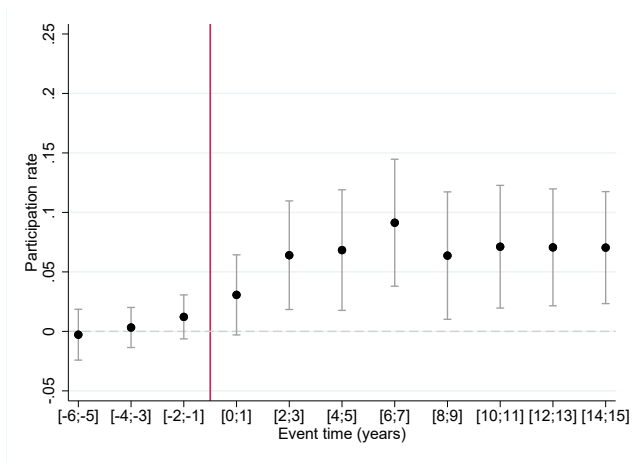
Placebo effect on participation



Note: Conditional on first predicted income bracket and not working in $t = -1$

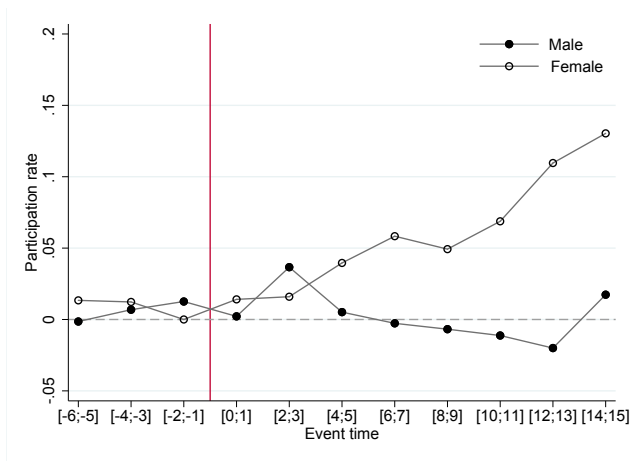
Back

Dynamics of participation response



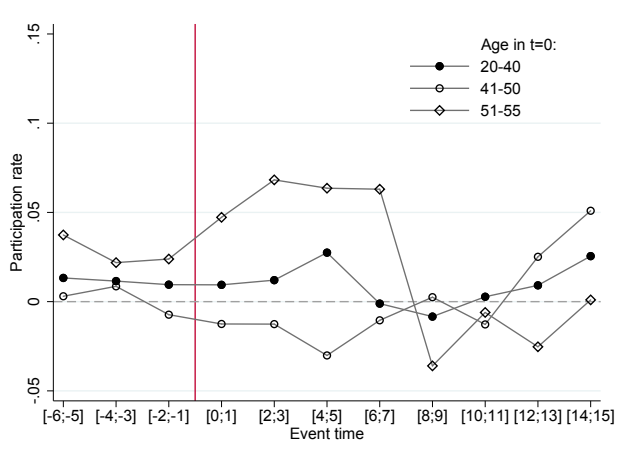
Back

Dynamics of participation response by gender



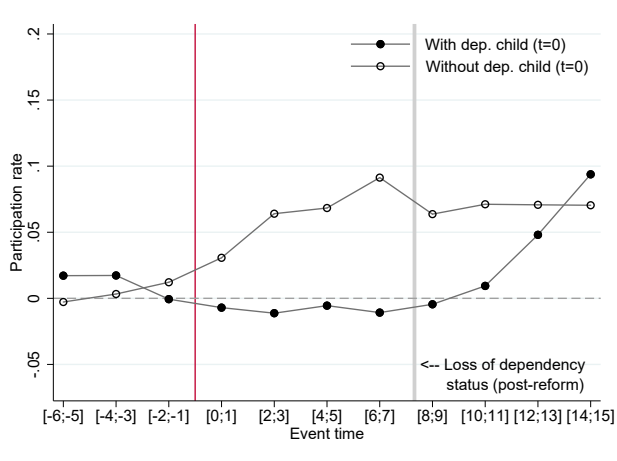
Back

Dynamics of participation response by age in $t = 0$



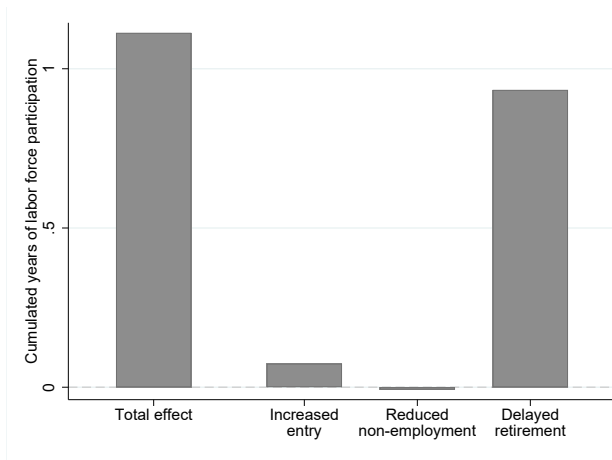
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Dynamics of participation response: dep. children in $t = 0$



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Decomposition of LFP response

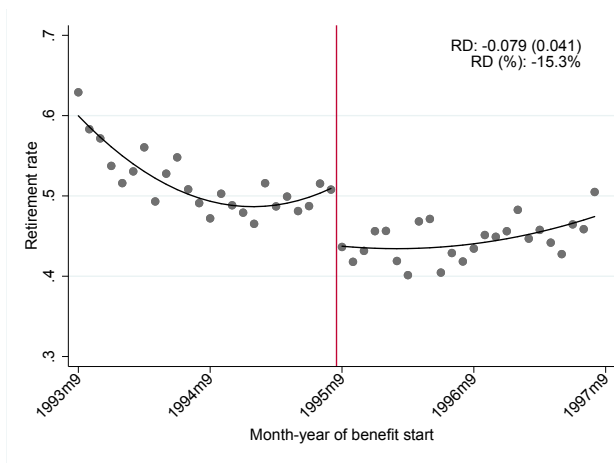


Retirement rate is 15.3% (8 p.p.) lower in $t = 15$

[Retirement rate](#)

[Back](#)

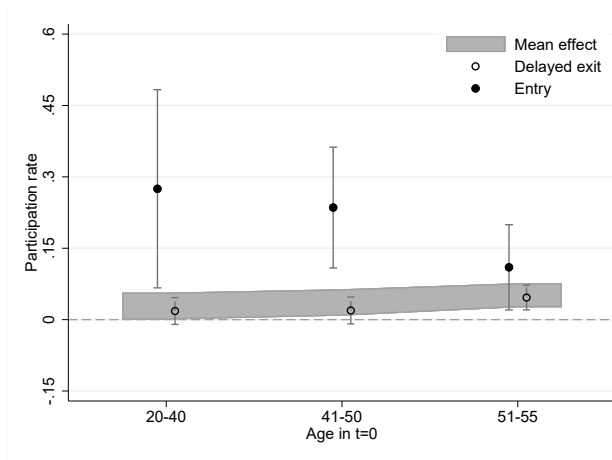
Retirement rate in $t = 15$



IV-RD ($\times 1000$): 0.100 (0.047) [Back](#)

Age profile of participation response

↓ exit at older ages, ↑ entry at young ages



Gender and age profile of labor supply response

	Gender		Age in $t = 0$		
	Female (1)	Male (2)	20-40 (3)	41-50 (4)	51-55 (5)
Participation rate	0.101*** (0.012) [0.639]	0.045*** (0.017) [0.553]	0.028** (0.014) [0.883]	0.036*** (0.014) [0.585]	0.051*** (0.013) [0.212]
Days worked	1.084 (2.641) [341.62]	1.084 (3.863) [338.91]	5.279* (2.792) [348.28]	-4.031 (3.502) [336.77]	5.926 (5.835) [326.57]
Daily wage	1.271 (1.430) [74.507]	-4.394* (2.509) [83.966]	1.049 (1.868) [73.886]	1.508 (1.884) [80.890]	-0.747 (3.412) [80.507]
Benefit-start-month FE	x	x	x	x	x
Calendar year FE	x	x	x	x	x
Linear trend	x	x	x	x	x
Quadratic trend	x	x	x	x	x

Heterogeneity of labor supply response by presence of dependent children in $t = 0$

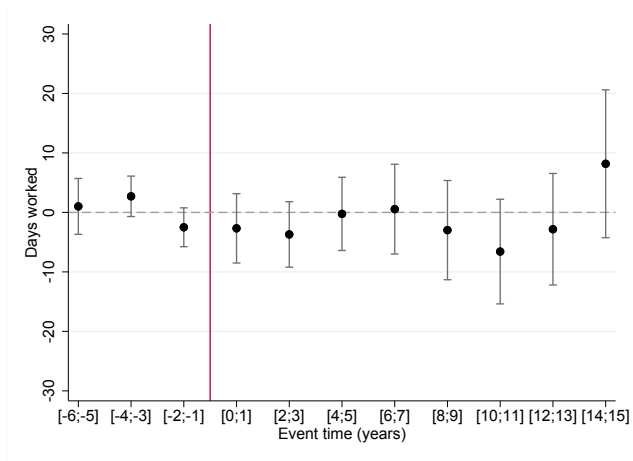
	With dependent children (1)	Without dependent children (2)
Participation rate	0.036*** (0.011) [0.719]	0.076*** (0.010) [0.457]
Days worked	-2.794** (1.317) [353.03]	-1.478 (1.587) [349.78]
Daily wage	-2.412** (1.226) [78.218]	0.252 (1.324) [74.908]
Benefit-start-month FE	x	x
Calendar year FE	x	x
Linear trend	x	x
Quadratic trend	x	x

Other margins of adjustment of labor supply

	Benefit ($\times 1,000$)	Control mean	Observations
Full-time job	0.010* (0.005)	0.891	68253
Change firm	-0.004 (0.005)	0.082	68253
Change industry	-0.002 (0.003)	0.029	68253
Change province	-0.000 (0.003)	0.025	68253

Note: Conditional on work experience in $t = 0$. Subsample of workers employed in private sector. [Back](#)

No *long-term* effect on intensive margin



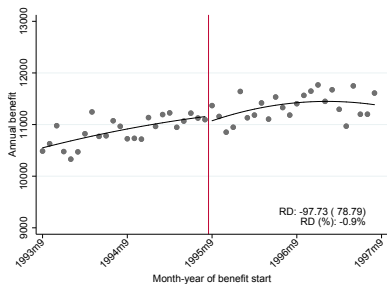
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Benefit replacement rate

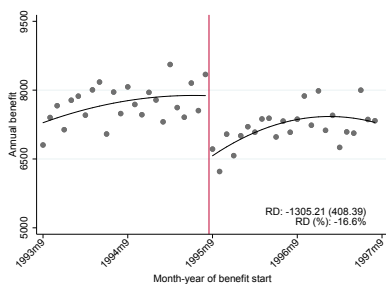
	Benefit start date	
	Before Sept 1, 1995	After Sept 1, 1995
<i>Spouse (with and without children)</i>		
Spouse only		
Survivor's taxable income $\leq 3 \times$ minimum pension	60%	60%
Survivor's taxable income $\leq 4 \times$ minimum pension	60%	45%
Survivor's taxable income $\leq 5 \times$ minimum pension	60%	36%
Survivor's taxable income $> 5 \times$ minimum pension	60%	30%
Spouse with one child	80%	80%
Spouse with two or more children	100%	100%

Benefit before and after loss of child dependency

With dependent children



Upon loss of child dependency



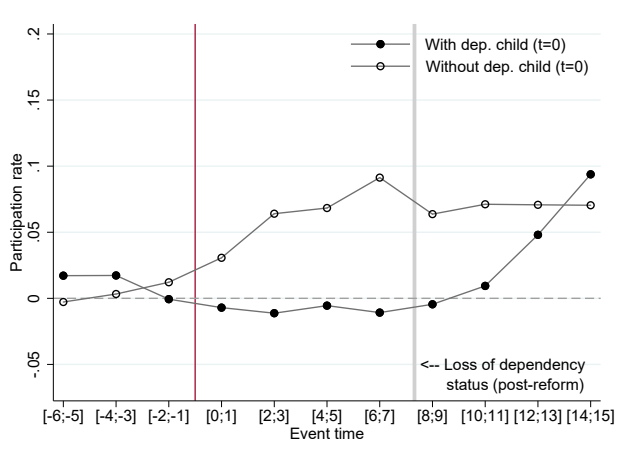
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Placebo

Number of years with dependent children	
<i>Placebo threshold</i>	
September 1992	-0.404 (0.568)
September 1993	0.757 (0.423)
September 1994	-1.317*** (0.413)
September 1995	1.223*** (0.415)
September 1996	-0.345 (0.421)
September 1997	0.390 (0.416)
September 1998	-0.502 (0.540)

Note: 24-month bandwidth [Back](#)

Dynamics of participation response: dep. children in $t = 0$



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A model of LFP decision

Model setup

- $\max u(c) - \mathbb{I}\{l = 1\} \phi$, $u(\cdot)$ concave
- s.t. $c = \mathbb{I}\{l = 1\} z + B$
- $\phi \sim F(\phi)$ is disutility of work, $F(\cdot)$ extreme value distribution $\rightarrow f'(\cdot) < 0$

Participation decision

- Let $V(z, l, B)$ be indirect utility function
- Work iff $V(z, 1, B) - V(0, 0, B) \geq \phi \iff \phi \leq \bar{\phi}(z, B)$
- LFP rate is $\Phi(z, B) = F(\bar{\phi}(z, B))$

High value or low cost?

LFP response to benefit change:

$$\frac{d\Phi}{dB} = - \underbrace{\frac{d\Phi}{dz}}_{\text{Cost of work (-)}} \cdot \underbrace{\gamma}_{\text{Risk aversion (Utility value (+))}} \cdot \frac{z}{B}$$

- $|d\Phi/dB|$ increasing in γ , i.e. if utility over consumption is strongly curved and $u'(c)$ rises sharply when B falls
- $|d\Phi/dB|$ increasing in $d\Phi/dz$, which is negative function of work disutility

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Derivation

- LFP rate $\Phi(z, B) = F(\bar{\phi}(z, B))$
- Income effect of B on LFP rate

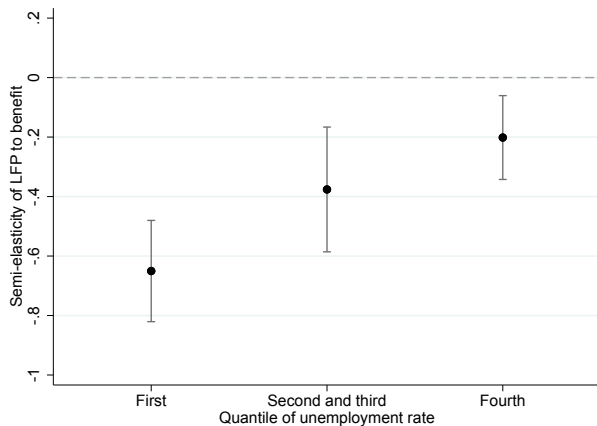
$$\frac{d\Phi}{dB} = f(\bar{\phi}) \cdot \left[\frac{\partial V(z, 1, B)}{\partial B} - \frac{\partial V(0, 0, B)}{\partial B} \right] \approx f(\bar{\phi}) \cdot u''(c(B)) \cdot z$$

- Elasticity of LFP to earnings $\frac{d\Phi}{dz} = f(\bar{\phi}) \cdot u'(c(B))$
- Sub in for $f(\bar{\phi})$

$$\frac{d\Phi}{dB} \approx \frac{d\Phi}{dz} \cdot \frac{u''(c(B))}{u'(c(B))} \cdot z = -\frac{d\Phi}{dz} \cdot \gamma \cdot \frac{z}{B}$$

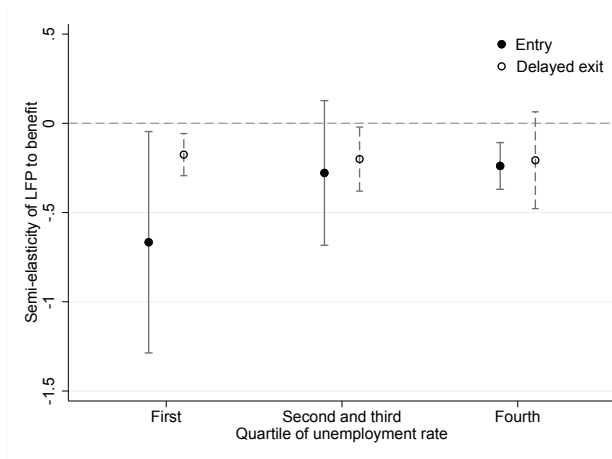
Heterogeneity wrt regional unemployment rate

Controlling for rate of undeclared work



Heterogeneity wrt regional unemployment rate

New entrants vs. incumbents



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Proof

Model setup

- $\max u(c) - \mathbb{I}\{l = 1\} \phi$
- s.t. $c = \mathbb{I}\{l = 1\} z + B$
- $\phi \sim F(\phi)$ is disutility of work

Participation decision

- Let $V(z, l, B)$ be indirect utility function
- Work iff $V(z, 1, B) - V(0, 0, B) \geq \phi \iff \phi \leq \bar{\phi}(z, B)$
- LFP rate is $\Phi(z, B) = F(\bar{\phi}(z, B))$

Proof (cont.)

Participation response and value of benefit

- Semi-elasticity of LFP to benefit

$$\begin{aligned}\frac{d\Phi}{d \log B} &= f(\bar{\phi}) \cdot \frac{\partial \bar{\phi}}{\partial B} \cdot B = f(\bar{\phi}) \cdot \left[\frac{\partial V(z, 1, B)}{\partial B} - \frac{\partial V(0, 0, B)}{\partial B} \right] \cdot B \\ &\approx f(\bar{\phi}) \cdot \frac{\partial^2 V}{\partial z \partial B} \cdot z \cdot B = f(\bar{\phi}) \cdot u''(c(B)) \cdot z \cdot B\end{aligned}$$

- Semi-elasticity of LFP to earnings $\varepsilon = \frac{d\Phi}{d \log z} = f(\bar{\phi}) \cdot u'(c(B)) \cdot z$
- Rescaling $d\Phi/d \log B$ by ε

$$\frac{\left[\frac{d\Phi}{d \log B} \right]}{\varepsilon} \approx \frac{u''(c(B)) \cdot B}{u'(c(B))} \approx \frac{u'(c(B)) - u'(c(0))}{u'(c(B))}$$

RP approach: applicability and limitations

- Application of Chetty [2008] and Landais [2015] to *within*-state valuation of unconditional transfer
- Robust to state-dependent utility
- Based on labor supply data and within-state policy variation
 - Wide applicability
 - Avoids limitations of consumption-based implementation approaches
- Assumption of optimizing behavior: no frictions, absence or separability of other margins of adjustment
- Can be extended to two-state setting to evaluate value of insurance